# Expectation-Maximization (EM)

Saturday, October 6, 2018 12:21

For the glory of God

#### Maximum Likelihood Estimation (ULE)

· As we know, Maximum Litelihood Estimation (ULE) attempts to find the parameter value that maximizes the litelihood

Given observed data. What is the chance that a given reality is true?

· Bostcally, what MLE does is to maximize 109-likelihood function to estimate the parameter value.

we need to take logarithm because maximization of the likelihood function is difficult in general

- · If we assume the distribution as Gaussian distribution, the log-likelihood function is as following;
  - a) For univariate Gaussian distribution

$$P(X \mid \mathcal{M}, \delta^2) = \frac{1}{\sqrt{2\pi\delta^2}} \exp\left(-\frac{1}{2} \frac{(X - \mathcal{M})^2}{2\delta^2}\right)$$
; where  $\delta^2 = Variance$ ,  $\mathcal{M} = mean$ 

b) For multi-variate Gaussian distribution \*\* 1 100 x

$$P\left(X\mid\mathcal{M},\Sigma\right) = \frac{1}{\left(\left(2\pi\right)^{0/3}\mid\Sigma\right)^{1/2}} \exp\left[-\frac{1}{2}\left(X-\mathcal{M}\right)^{\mathsf{T}}\Sigma^{-1}\left(X-\mathcal{M}\right)\right] \quad \text{; Where } \Sigma = \text{Covariance, } D = \text{Dimension}$$

- · In the end, MLE will plovide the best parameters (M, 62, 2) that fit well given data (X).
- Then, how does ULE find the best parameters?
  - The basic idea is to take the derivative and to equate it to zero.

$$\frac{JP(x|\theta)}{J\theta}=0 \rightarrow \theta_{MLE}$$
; where  $\theta$  represents parameters  $(\mu, b^2, ...)$ 

$$\Leftrightarrow \Theta_{ME} = \text{arg max } P(X|\theta)$$

· For example, let's say that we have N number of data potents. (X1, X2, ..., XN)

= argmax 
$$\pi$$
 P(x718) ; Here, we assumed the data points are IID (Independently Identically Distributed)

 $\theta$ 

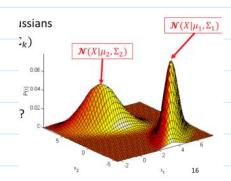
= 
$$\underset{\Delta}{\text{arg max}} \left( \frac{1}{\sqrt{2\pi G^2}} \right)^N \exp \left( -\frac{1}{2G^2} \sum_{i=1}^N (\chi_i - \mu_i)^2 \right)$$
; if p is univertiate Gaussian distribution with the older points

#### Gaussian Utixture Model (GMM)

The probability (P) may be multi-modal as a mixture of uni-modal distributions.

$$P(X) = \sum_{k=1}^{K} \pi_k N(X | \mathcal{U}_k, \mathcal{Z}_k)$$
 ; where K is number of Gaussians

Wixing coefficient: Weightage for each Gaussian distribution;  $0 \le \pi k \le 1$  and  $\sum_{n=1}^{K} \pi k = 1$ 



: two Gaussian distributions

e.g. T1 = 6.3 / T2 = 0.7

- · Let's say that we want to maximize likelihood of the multi-modal distribution.
- In a similar way, we would take logarithm and defivative;

$$ln P(X) = \sum_{N=1}^{N} ln \left( \sum_{k=1}^{K} \pi_{k} N(X_{k} | \mathcal{M}_{k}, \Sigma_{k}) \right)$$

 $\frac{1}{10}$  In  $P(x|0) = ? \Rightarrow MLE$  doesn't work because there is no closed form solution

(Non-convex difficult)

· Therefore, for this case, parameters can be calculated using EM algorithm.

# Expectation - Maximization (EM) Algorithm based on GMM

## Introduction

- The EM algorithm was explained and given its name in 1977.
- · The EM algorithm has been used to find ULE of parameters in a statistical model (e.g. Gaussian) where the equation can't be solved directly
- Typically, the model involves latent variables (Hiddew variable)
- · Thus, I would say that;
- The EM algorithm is an efficient iterative procedure to compute the MLE in the presence of hidden data.

## Latent Variable

- · Statistically, ladent variables are variables that council be observed directly but its values con be infleted by taking other measurements.
  - e.g. we may not be able to cliectly quantify intelligence but we think it exists.
    - 4 Your intelligence is a latent unitable that affects your performance on multiple tasks even if it can't be directly measured.
- · When we think about the probability,
- We can think of the mixing coefficients as prior probabilities.  $P(k) = \pi_k$

- For a given value of x, we can evaluate the corresponding posterior probabilities.  $h_k(x) = P(k|x)$ 

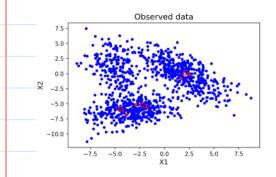
4 This is called as responsibilities  $t_k(x)$ . Then t = |alent variable|

$$p(k|x) = \frac{p(k)p(x|k)}{p(x)}$$
 ; Based on Bayes tule

$$=\frac{\mathcal{T}_{K} N\left(X \mid \mathcal{A}_{K}, \mathcal{Z}_{K}\right)}{\sum\limits_{\substack{\Gamma = 1 \\ \Gamma = 1}} \mathcal{T}_{\Gamma} N\left(X \mid \mathcal{A}_{\Gamma}, \mathcal{Z}_{\Gamma}\right)}; \text{ where } \mathcal{T}_{K} = \frac{\mathcal{N}_{K}}{\mathcal{N}} \Rightarrow \mathcal{N}_{K} = \text{ number of data powds assigned to cluster } K.$$

## How EM algorithm works

- Step 1) Initialization
- Let's assume that K=3.
- As the first step, we will place three pourts randomly on the observed data



## Python code

[N,d] = X.shape
K = 3
C = np.zeros([K,d])
for i in range(K):
 random\_index = np.random.randint(N)
 The potots are on the observed data.

; where X = data set, random index would give you K random points

- As the second step, we will initialize the values of Ui, Zi, and Ii.
  - ; For mixing coefficient ( $\pi_1$ ),  $\pi_1 = 1/k$  :  $\pi_1 = 0.3$ ,  $\pi_2 = 0.3$ ,  $\pi_3 = 0.3$
  - ; For mean  $(\mathcal{U}_{\bar{1}}), \ \mathcal{U}_{\bar{1}} = \text{CETJ}$  ; it means the control are considered as a mean.

e.g. 
$$C_1 = \mathcal{U}_1 = [-3.16, -6.45]$$

; For covariance ( $\Sigma_{\bar{1}}$ ),  $\Sigma_{\bar{1}}$  = numpy. eye(d) ; where d is dimension

e.g. 
$$\Sigma_{1} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$
,  $\Sigma_{2} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$ ,  $\Sigma_{3} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$ 

$$= \begin{bmatrix} 1 & \text{sigma} \\ \text{[array([[1., 0.], [0., 1.]]), array([[1., 0.], (0.], [0., 1.]]), array([[1., 0.], (0.], [0., 1.]]))} \\ \begin{bmatrix} 0 & 1 & 1 \\ 0 & 1 & 1 \end{bmatrix}$$

1 sigma = [np.eye(d)]\*K

## Step 2) Expectation

- As the first step, evaluate the responsibilities using the cuteut parameter values.
- As the Second slep, Evaluate the indicators for each data point.
- Then, we will see if the data point gets involved in K=1, K=2, or K=3.
- Let us take an example. Let's suppose that we have X (1000,2) daylaset.

$$X = \begin{bmatrix} 0.1, -0.2 \\ \vdots \\ 0.6, -0.4 \end{bmatrix}$$
 j  $1000 \times 2$  moltrix

, Let's calculate Gaussian PDF for K=1 cose first.

$$\pi_1 = 0.3$$
,  $\mathcal{M}_1 = \begin{bmatrix} -3.16, -6.45 \end{bmatrix}$ ,  $\Sigma_1 = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$ 

$$P\left(\left.X_{1}\right|\mathcal{M},\Sigma\right) = \frac{1}{\left(\left(2\pi\right)^{0/2}\left|\Sigma\right|\right)^{1/2}} \exp\left[-\frac{1}{2}\left(\left.X_{1}\right|\mathcal{M}\right)^{T} \mathcal{Z}^{-1}\left(\left.X_{1}\right|\mathcal{M}\right)\right] \quad ; \quad \text{where } D=2$$

$$\Rightarrow P(X_{1} | \mathcal{L}_{1}, \Sigma_{1}) = \begin{bmatrix} 1.9 \times 10^{-14} \\ \vdots \\ 3.1 \times 10^{-1} \end{bmatrix} \quad ; \quad |000 \times 1| \quad |000 \times 1|$$

Then,

$$\uparrow_{1}(X) = \frac{\mathcal{I}_{1} \times P(X_{\overline{J}} | \mathcal{M}_{1}, \overline{\Sigma}_{1})}{\sum_{i=1}^{K} \mathcal{I}_{1} P(X_{\overline{J}} | \mathcal{M}_{1}, \overline{\Sigma}_{1})} = \frac{1000 \times 1 \quad \text{modifix}}{1000 \times 1 \quad \text{modifix}} = \begin{bmatrix} 6.35 \times 10^{-15} \\ \vdots \\ 3.3 \times 10^{-1} \end{bmatrix}$$

$$r_3(x) = |000x| \text{ modifix}$$

Hence,

$$+ = \begin{bmatrix} \vdots & \vdots & \vdots & \vdots \\ \vdots & \vdots & \vdots & \vdots \end{bmatrix} = \begin{bmatrix} h_1, h_2, h_3 \end{bmatrix}$$
 is loop x 3 moltrix

- For the next step, we're going to estimate the indicator for each data powd

Literally, the code will ask to each point 'Does it look like it came from k=1, k=2, or k=3?'

Soy, we have the first you of 
$$1 = [0.1, 0.2, 0.3]$$

4 Indicator will select the third column because of argmax.

## Calculate argmax of the responsibility and choose the indicator for each data point properly
for i in range(N):
 I[i] = r[i].argmax()+1 # +1 is needed because of Pythonic characteristic

- Finally,  $I = \begin{bmatrix} 1 \\ 2 \\ 3 \end{bmatrix}$  telling us the first data point might come from K=1 clustering
- Step 3 ) Maximizaction
- Re-estimate the parameters using the current responsibilities
- In terms of π,

$$\mathcal{I}_{J} = \frac{1}{N} \sum_{n=1}^{N} Y_{J}(X_{n})$$
 ; where  $N = \text{number of data}$  point,  $J = \text{number of dustering}$ 

e.g. 
$$\pi_1 = \frac{1}{1000} \left( h_1(X_1) + h_1(X_2) + h_1(X_3) + \dots + h_1(X_{1000}) \right) = \text{Constant}$$

- In terms of M.

$$\mathcal{M}_{2} = \frac{\sum_{n=1}^{M-1} f_{2}(\chi_{N}) \cdot \chi_{N}}{\sum_{n=1}^{N} f_{2}(\chi_{N}) \cdot \chi_{N}}$$

e.g. 
$$\mathcal{U}_1 = \frac{A}{\text{constant}} = [X_1, X_2]$$
; where  $A = F_1(X_1) \cdot (X_1, X_2)_1 + F_1(X_2) \cdot (X_1, X_2)_2 + \cdots$ 

- In terms of Σ,

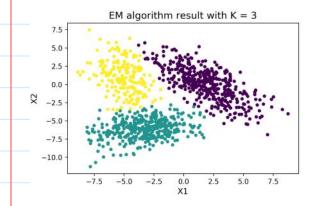
$$Z_{\bar{J}} = \frac{\sum_{j=1}^{N} Y_{j}(X_{ij}) \cdot (X_{ij} - M_{\bar{J}})}{\sum_{j=1}^{N} Y_{j}(X_{ij})}$$
; where  $M_{\bar{J}} = \text{new mean (or Centrolid)}$ 

e.g. 
$$\Sigma_1 = \frac{A}{\text{constant}} = \begin{bmatrix} 2 \times 2 \end{bmatrix}$$
 is where  $A = h_1(X_1) \cdot (X_1, X_2)_1^T \cdot (X_1, X_2)_1 + \cdots$ 

2×1 1×2

- So, it feels like that we have new parameters.

- · Step 4) Repeat the process 2~3 with Convergence criteria is satisfied.
- We will end up gelting a kind of result as below;



#### Examples

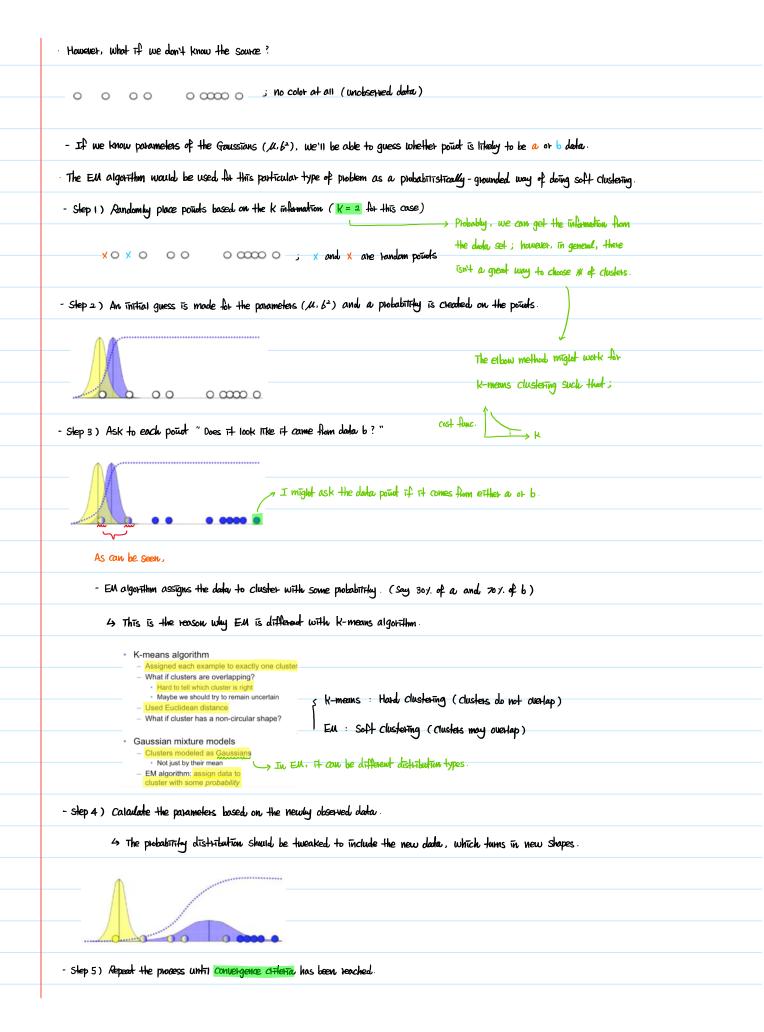
One - Dimensional

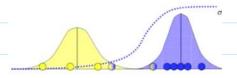
Let's say we have the following observed dadau;

- It seems like there one K=2 Gaussians with Lunknown 11 and 62
- · In the end, we will get the result; It can be caladaded by; for b data,

servation 
$$\mathcal{U}_b = \frac{\chi_1 + \chi_2 + \dots + \chi_{N-b}}{\chi_b}; \text{ where } N = \text{Number of close}$$

$$b_b^2 = \frac{(x_1 - x_b)^2 + \dots + (x_n - x_b)^2}{n_b}$$





; It might be slow on large data set.

; Hence, it may reveal hidden patherns. Note that it won't work well it clusters contain the points.

# Two-dimensional

